

ON A BINOMIAL COEFFICIENT AND A PRODUCT OF PRIME NUMBERS

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Let p_n be the n -th prime number. We prove the following double-inequality. For all integers $k \geq 5$ we have

$$\exp[k(c_0 - \log \log k)] \leq \frac{\binom{k^2}{k}}{p_1 \cdot p_2 \cdots p_k} \leq \exp[k(c_1 - \log \log k)]$$

with the best possible constants

$$c_0 = \frac{1}{5} \log 23 + \log \log 5 = 1.10298 \dots$$

and

$$c_1 = \frac{1}{192} \log \binom{36864}{192} + \log \log 192 - \frac{1}{192} \log (p_1 \cdot p_2 \cdots p_{192}) = 2.04287 \dots$$

This refines a result published by GUPTA and KHARE in 1977.

1. INTRODUCTION

The work on this note has been inspired by a remarkable short paper published by GUPTA and KHARE [4] in 1977. The authors presented a connection between the binomial coefficient $\binom{k^2}{k}$ and the product of the first k prime numbers.

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Proposition. *If $k = 3, 4, \dots, 1793$, then*

$$\binom{k^2}{k} > p_1 \cdot p_2 \cdots p_k,$$

whereas, if $k \geq 1794$, then

$$\binom{k^2}{k} < p_1 \cdot p_2 \cdots p_k.$$

The Proposition implies an interesting number theoretical theorem: if $1794 \leq k \leq n \leq k^2$, then $\binom{n}{k}$ has less than k distinct prime divisors. This improves an earlier result of ERDÖS, GUPTA and KHARE [3].

From the Proposition we obtain

$$(1.1) \quad Q_k = \frac{\binom{k^2}{k}}{p_1 \cdot p_2 \cdots p_k} < 1 \quad \text{for } k \geq 1794.$$

It is natural to look for a refinement of (1.1). More precisely, we ask for the largest number c_0 and the smallest number c_1 such that the double-inequality

$$(1.2) \quad \exp[k(c_0 - \log \log k)] \leq Q_k \leq \exp[k(c_1 - \log \log k)]$$

holds for all $k \geq 5$. It is our aim to solve this problem. In the next section, we demonstrate that (1.2) is valid with $c_0 = 1.10298\dots$ and $c_1 = 2.04287\dots$. This provides not only a positive lower bound for Q_k , but improves (1.1) for all $k \geq 2237$.

2. MAIN RESULT

In order to offer sharp bounds for Q_k we need the following estimates.

Lemma 1. *For $k \geq 2$ we have*

$$(2.1) \quad k^2 \log k - \left(k^2 - k + \frac{1}{2}\right) \log(k-1) - \frac{1}{2} \log(2\pi) - \frac{1}{12k} - \frac{1}{12(k-1)k^2} \\ < \log \binom{k^2}{k} < k^2 \log k - \left(k^2 - k + \frac{1}{2}\right) \log(k-1) - \frac{1}{2} \log(2\pi).$$

This lemma is due to SASVÁRI [9]. Moreover, we need upper and lower bounds for the Chebyshev function

$$\theta(x) = \sum_{p \leq x} \log p,$$

where p runs over all prime numbers $\leq x$.

Lemma 2. *We have for $k \geq 3$*

$$(2.2) \quad \frac{1}{k}\theta(p_k) = \frac{1}{k} \sum_{j=1}^k \log(p_j) \geq \log k + \log \log k - 1 + \frac{\log \log k - 2.1454}{\log k}$$

and for $k \geq 126$

$$(2.3) \quad \frac{1}{k}\theta(p_k) \leq \log k + \log \log k - 1 + \frac{\log \log k - 1.9185}{\log k}.$$

A proof of Lemma 2 can be found in [7]; see also [2]. We are now in a position to present our main result.

Theorem. *For all integers $k \geq 5$ we have*

$$(2.4) \quad \exp[k(c_0 - \log \log k)] \leq \frac{\binom{k^2}{k}}{p_1 \cdot p_2 \cdots p_k} \leq \exp[k(c_1 - \log \log k)]$$

with the best possible constants

$$(2.5) \quad c_0 = \frac{1}{5} \log 23 + \log \log 5 = 1.10298 \dots$$

and

$$(2.6) \quad c_1 = \frac{1}{192} \log \binom{36864}{192} + \log \log 192 - \frac{1}{192} \log(p_1 \cdot p_2 \cdots p_{192}) = 2.04287 \dots$$

Proof. A short calculation reveals that (2.4) is equivalent to

$$c_0 \leq f(k) \leq c_1$$

with

$$(2.7) \quad f(k) = \frac{1}{k} \log \binom{k^2}{k} + \log \log k - \frac{1}{k}\theta(p_k).$$

First, we prove that

$$(2.8) \quad 1.9114 < f(k) \quad \text{for } k \geq 126.$$

Let $k \geq 126$. The left-hand side of (2.1) and the elementary inequalities

$$\frac{1}{k} < \log k - \log(k-1) < \frac{1}{k-1}$$

imply

$$(2.9) \quad \log \binom{k^2}{k} > k + \left(k - \frac{1}{2}\right) \log(k-1) - \alpha_0 > k + \left(k - \frac{1}{2}\right) \log k - \alpha_1$$

with

$$\alpha_0 = \frac{1}{2} \log(2\pi) + \frac{1}{12 \cdot 126} + \frac{1}{12 \cdot 125 \cdot 126 \cdot 126}$$

and

$$\alpha_1 = \alpha_0 + 1 + \frac{1}{2 \cdot 125} = 1.9235 \dots$$

Using (2.3), (2.7), and (2.9) leads to

$$f(k) > 2 - \left(\frac{\log k}{2k} + \frac{1.9236}{k} + \frac{\log \log k - 1.9185}{\log k} \right).$$

Since

$$\frac{\log k}{2k} + \frac{1.9236}{k} < 0.0345 \quad \text{and} \quad \frac{\log \log k - 1.9185}{\log k} < 0.0541,$$

we get

$$f(k) > 2 - 0.0345 - 0.0541 = 1.9114 \dots$$

This settles (2.8). By direct computation we find

$$(2.10) \quad \min_{5 \leq k \leq 125} f(k) = f(5) = 1.10298 \dots$$

From (2.8) and (2.10) we conclude that the first inequality in (2.4) holds for $k \geq 5$ with the best possible constant given in (2.5).

Applying the second inequality in (2.1), (2.2), and (2.7) yields for $k \geq 652$:

$$(2.11) \quad f(k) < 1 + (k-1) \log \left(1 + \frac{1}{k-1} \right) - \frac{\log(2\pi(k-1))}{2k} - \frac{\log \log k - 2.1454}{\log k} \\ < 2 - \frac{\log \log k - 2.1454}{\log k} < 2.0428.$$

Moreover, we have

$$(2.12) \quad \max_{5 \leq k \leq 651} f(k) = f(192) = 2.04287 \dots$$

From (2.11) and (2.12) we obtain that the right-hand side of (2.4) is valid with the best possible constant c_1 given in (2.6). \square

3. REMARKS

(i) The behaviour of $f(k)$ for large k is quite surprising. We have the limit relation

$$\lim_{k \rightarrow \infty} f(k) = 2,$$

although computer calculations show that $f(k)$ is decreasing in the range $1000 \leq k \leq 1\,500\,000$ with $f(10\,000) = 1.9908 \dots$

(ii) More inequalities involving the product $p_1 \cdot p_2 \cdots p_k$ can be found in [1], [5, p. 246], [6].

(iii) Let $\psi = \Gamma'/\Gamma$ be the logarithmic derivative of Euler's gamma function. We define for $x \geq 2$:

$$\phi(x) = \log \frac{\Gamma(x^2 + 1)}{\Gamma(x+1)\Gamma(x^2 - x + 1)}.$$

Differentiation gives

$$\phi'(x) = 2x[\psi(x^2 + 1) - \psi(x^2 - x + 1)] + [\psi(x^2 - x + 1) - \psi(x + 1)].$$

Since ψ is strictly increasing on $(0, \infty)$, we conclude that $\phi'(x)$ is positive for $x \geq 2$. Hence,

$$\phi(k) = \log \binom{k^2}{k} < \log \binom{(k+1)^2}{k+1} = \phi(k+1) \quad \text{for } k \geq 2.$$

We have $\phi(1) = 0 < \log 6 = \phi(2)$. Thus we obtain: If $\mu \geq 0$, then the sequence

$$\Delta_k(\mu) = \binom{k^2}{k} \exp(\mu p_k)$$

is strictly increasing for $k \geq 1$. This result leads to the question: Does there exist a negative real number μ_0 and an integer k_0 such that $\Delta_k(\mu_0)$ is increasing for $k \geq k_0$? We show that the answer is "no".

We assume that

$$\Delta_k(\mu_0) \leq \Delta_{k+1}(\mu_0) \quad \text{for } k \geq k_0.$$

This is equivalent to, say

$$(3.1) \quad -\mu_0 \leq \frac{\log \binom{(k+1)^2}{k+1} - \log \binom{k^2}{k}}{p_{k+1} - p_k} = \sigma_k.$$

Using the asymptotic formula

$$\log \binom{k^2}{k} = \left(k - \frac{1}{2}\right) \log k + k - \frac{1}{2}(1 + \log(2\pi)) + O(1/k),$$

see [4], we get

$$(3.2) \quad \log \binom{(k+1)^2}{k+1} - \log \binom{k^2}{k} = \log k + 2 + O(1/k).$$

Applying (3.2) and the limit relations

$$\lim_{k \rightarrow \infty} \frac{\log p_k}{\log k} = 1 \quad \text{and} \quad \liminf_{k \rightarrow \infty} \frac{\log p_k}{p_{k+1} - p_k} = 0,$$

see [8], [10], yields

$$(3.3) \quad \liminf_{k \rightarrow \infty} \sigma_k = \liminf_{k \rightarrow \infty} \left(\frac{\log p_k}{p_{k+1} - p_k} \cdot \frac{\log k + 2 + O(1/k)}{\log p_k} \right) = 0.$$

From (3.1) and (3.3) we conclude that $\mu_0 \geq 0$.

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